

## EDGEWORTH EXPANSION

Consider the central limit theorem: For  $X_1, \dots, X_n$  iid with mean  $\mu$  and variance  $\sigma^2 < \infty$ ,

$$\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \rightarrow N(0,1) \text{ in dist. as } n \rightarrow \infty.$$

We might wish to know

- How fast is the convergence to Normality?
- What features of the distribution of  $X_1, \dots, X_n$  affect the rate of convergence and how?

Edgeworth expansions help us to answer these questions. Moreover, they can be used to show that the CLT works — and that it can provide a better approximation to the sampling distribution of  $\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma}$  than the  $N(0,1)$  distribution.

Theorem (1<sup>st</sup> and 2<sup>nd</sup> order Edgeworth Expansions):

Let  $Y_1, \dots, Y_n$  be iid with  $\mathbb{E} Y_i = \mu$ ,  $\text{Var} Y_i = \sigma^2 \in (0, \infty)$ ,  $\mathbb{E} |Y_i|^3 < \infty$ ,  $\mathbb{E} |Y_i|^4 < \infty$ , and suppose

$$|\mathbb{E} \exp(itY_i)| < 1 \text{ for all } t \neq 0. \text{ (Non-lattice)}$$

Then

$$\sup_{x \in \mathbb{R}} \left| P\left(\frac{\sqrt{n}(\bar{Y}_n - \mu)}{\sigma} \leq x\right) - \Phi_{n,3}(x) \right| = o(n^{-1/2})$$

as  $n \rightarrow \infty$  and

$$\sup_{x \in \mathbb{R}} \left| P\left(\frac{\sqrt{n}(\bar{Y}_n - \mu)}{\sigma} \leq x\right) - \Phi_{n,4}(x) \right| = o(n^{-1}),$$

where

$$\Psi_{n,3}(x) = \Phi(x) - \frac{1}{6\sqrt{n}} \frac{\mu_3}{\sigma^3} (x^2-1) \phi(x)$$

$$\Psi_{n,4}(x) = \Phi(x) - \left\{ \frac{1}{6\sqrt{n}} \frac{\mu_3}{\sigma^3} (x^2-1) + \frac{1}{24n} \cdot \left( \frac{\mu_4}{\sigma^4} - 3 \right) (x^2-3x) + \frac{1}{72n} \cdot \frac{\mu_2^2}{\sigma^6} (x^5 - 10x^3 + 15x) \right\} \phi(x).$$

We now spend some time deriving the E.E.: Here I acknowledge my indebtedness to Dr. David Hunter at Penn State, whose notes I have essentially followed.

We will need to introduce Hermite polynomials:

Hermite polynomials: The Hermite polynomials  $H_1, H_2, \dots$  are defined by the relation

$$(-1)^k \frac{d^k}{dx^k} \phi(x) = H_k(x) \phi(x), \quad \text{for } k=1,2,\dots$$

*b/c  $\frac{d}{dx} \phi(x)$  always brings a minus sign down to the front.*

We obtain, for  $k=1,2,3$ , the following:

$$\begin{aligned} \underline{k=1}: \quad \frac{d}{dx} \phi(x) &= \frac{d}{dx} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \\ &= -x \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \\ &= (-1) \underbrace{x}_{H_1(x)} \phi(x) \end{aligned}$$

$$\begin{aligned} \frac{d}{dx} [H_k(x) \phi(x)] &= (-1)^k \frac{d^{k+1}}{dx^{k+1}} \phi(x) \\ &= (-1)^{k+1} H_{k+1}(x) \phi(x) \end{aligned}$$

so  $H_1(x) = x$  since  $(-1)^1 \phi'(x) = H_1(x) \phi(x)$

$$\begin{aligned} \underline{k=2}: \quad \frac{d^2}{dx^2} \phi(x) &= \frac{d}{dx} [-x \phi(x)] \\ &= -x \phi'(x) - \phi(x) \end{aligned}$$

$$= x^2 \phi(x) - \phi(x)$$

$$= (x^2 - 1) \phi(x).$$

$$\text{So } H_2(x) = x^2 - 1, \text{ since } (-1)^2 \frac{d^2}{dx^2} \phi(x) = (x^2 - 1) \phi(x).$$

$$\begin{aligned} \underline{\underline{k=3}}: \quad \frac{d^3}{dx^3} \phi(x) &= \frac{d}{dx} (x^2 - 1) \phi(x) \\ &= (x^2 - 1) \phi'(x) + 2x \phi(x) \\ &= (x^2 - 1)(-1)x \phi(x) \\ &= (-1)[x^3 - x - 2x] \phi(x) \\ &= (-1) \underbrace{[x^3 - 3x]}_{H_3(x)} \phi(x). \end{aligned}$$

$$\text{So } H_3(x) = x^3 - 3x, \text{ since } (-1)^3 \frac{d^3}{dx^3} \phi(x) = (x^3 - 3x) \phi(x).$$

Doing more work, we can obtain

$$H_4(x) = x^4 - 6x^2 + 3$$

$$H_5(x) = x^5 - 10x^3 + 15x.$$

Note that the Hermite polynomials appear in the Edgeworth Expansions!

Next, we will use the fact that if  $X$  is a rv with characteristic function  $\psi_x$  satisfying  $\int_{-\infty}^{\infty} |\psi_x(t)| dt < \infty$ , then the density of  $X$  is given by

$$f_x(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} \exp(-itx) \psi_x(t) dt, \text{ for all } x \in \mathbb{R} \quad \left( \text{This is called the inversion formula} \right)$$

to obtain the following useful identity:

Useful Identity:

$$\begin{aligned} \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-itx} e^{-t^2/2} (it)^k dt &= \frac{(-1)^k}{2\pi} \int_{-\infty}^{\infty} \frac{d^k}{dx^k} e^{-itx} e^{-t^2/2} dt \\ &= (-1)^k \frac{d^k}{dx^k} \underbrace{\frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-itx} e^{-t^2/2} dt}_{= \phi(x) \text{ by the inversion formula, since } \exp(-t^2/2) \text{ is the characteristic function of the } N(0,1) \text{ dist.}} \\ &= (-1)^k \frac{d^k}{dx^k} \phi(x) \\ &= H_k(x) \phi(x). \end{aligned}$$

Derivation of Edgeworth expansions:

We will simplify slightly by assuming, without loss of generality, that  $X_1, \dots, X_n$  are iid with

$$\begin{aligned} \mathbb{E} X_1 &= 0 & \mathbb{E} X_1^3 &= 0 \\ \mathbb{E} X_1^2 &= 1 & \mathbb{E} X_1^4 &= \tau < \infty \end{aligned}$$

and that

$$|\mathbb{E} \exp(itX_1)| < 1 \quad \text{for all } t \neq 0.$$



We study the distribution of the standardized sum

$$S_n = \frac{1}{\sqrt{n}} \sum_{i=1}^n X_i = \sqrt{n} \left( \frac{\sum_{i=1}^n X_i}{n} - 0 \right).$$

Begin by constructing the characteristic function of  $S_n$ :

$$\begin{aligned} \psi_{S_n}(t) &= \mathbb{E} \exp(it S_n) \\ &= \mathbb{E} \exp\left(it \frac{1}{\sqrt{n}} \sum_{i=1}^n X_i\right) \\ &= \mathbb{E} \prod_{i=1}^n \exp\left(\frac{it}{\sqrt{n}} X_i\right) \\ &\stackrel{\text{(ind)}}{=} \prod_{i=1}^n \mathbb{E} \exp\left(\frac{it}{\sqrt{n}} X_i\right) \\ &= \left[ \mathbb{E} \exp\left(\frac{it}{\sqrt{n}} X_1\right) \right]^n \end{aligned}$$

The non-lattice condition comes in here. We need  $\mathbb{E} \exp(it X/\sqrt{n}) < 1$  so that we can raise it to the power  $n$  without making it diverge.

Now make a Taylor expansion of  $\mathbb{E} \exp\left(\frac{it}{\sqrt{n}} X\right)$  around  $t=0$ :

$$\begin{aligned} \mathbb{E} \exp\left[\frac{it}{\sqrt{n}} X\right] &= \mathbb{E} \left( \left[ 1 + \frac{itX}{\sqrt{n}} + \frac{1}{2} \left(\frac{itX}{\sqrt{n}}\right)^2 + \frac{1}{6} \left(\frac{itX}{\sqrt{n}}\right)^3 + \frac{1}{24} \left(\frac{itX}{\sqrt{n}}\right)^4 + \dots \right] + o(n^{-2}) \right) \\ &= \left( 1 + \frac{it^2}{2n} \right) + \frac{i}{6} \frac{(it)^3}{n^{3/2}} + \frac{i}{24} \frac{(it)^4}{n^2} + o(n^{-2}) \end{aligned}$$

$$\left[ e^x = 1 + x + \frac{x^2}{2} + \frac{x^3}{3!} + \dots \right]$$

□

Now we want to raise this to the  $n^{\text{th}}$  power:

$$\left( \mathbb{E} \exp \left[ \frac{zt}{\sqrt{n}} X \right] \right)^n = \left[ \left( 1 - \frac{t^2}{2n} \right) + \frac{\delta}{6} \frac{(zt)^3}{n^{3/2}} + \frac{\sigma}{24} \frac{(zt)^4}{n^2} + o(n^{-2}) \right]^n$$

$$= \left( 1 - \frac{t^2}{2n} \right)^n + n \left( 1 - \frac{t^2}{2n} \right)^{n-1} \frac{\delta}{6} \frac{(zt)^3}{n^{3/2}} + n \left( 1 - \frac{t^2}{2n} \right)^{n-1} \frac{\sigma}{24} \frac{(zt)^4}{n^2} + \frac{n(n-1)}{2} \left( 1 - \frac{t^2}{2n} \right)^{n-2} \left[ \frac{\delta}{6} \frac{(zt)^3}{n^{3/2}} \right]^2 + o(n^{-1})$$

$\frac{n!}{(n-1)! \cdot 1} = n$   
 $\frac{n!}{(n-2)! \cdot 2!}$

$$= \left( 1 - \frac{t^2}{2n} \right)^n + \left( 1 - \frac{t^2}{2n} \right)^{n-1} \left[ \frac{\delta}{6} \frac{(zt)^3}{n^{3/2}} + \frac{\sigma}{24} \frac{(zt)^4}{n^2} \right] + \left( 1 - \frac{t^2}{2n} \right)^{n-2} \frac{(n-1)}{n^2} \frac{\delta^2}{72} \frac{(zt)^6}{n^3} + o(n^{-1}),$$

where the second equality comes from the Multinomial Theorem

$$(a_1 + \dots + a_m)^n = \sum_{\substack{n_1, \dots, n_m \in \{0, \dots, n\}, \\ n_1 + \dots + n_m = n}} \binom{n!}{n_1! \dots n_m!} a_1^{n_1} \dots a_m^{n_m}.$$

□

Next, we use the fact that for each nonnegative integer  $k$ ,

$$(*) \quad \left(1 + \frac{a}{n}\right)^{n-k} = e^a \left(1 - \frac{a(a+k)}{2n}\right) + o\left(\frac{1}{n}\right) \quad \text{as } n \rightarrow \infty$$

to write

$$\left(1 - \frac{t^2}{2n}\right)^n = e^{-t^2/2} \left(1 + \frac{t^2/2(t^2/2)}{2n}\right) + o\left(\frac{1}{n}\right)$$

$$\left(1 - \frac{t^2}{2n}\right)^{n-1} = e^{-t^2/2} \left(1 + \frac{t^2/2(1-t^2/2)}{2n}\right) + o\left(\frac{1}{n}\right)$$

$$\left(1 - \frac{t^2}{2n}\right)^{n-2} = e^{-t^2/2} \left(1 + \frac{t^2/2(2-t^2/2)}{2n}\right) + o\left(\frac{1}{n}\right).$$

Plugging these into our expression for  $\left(\mathbb{E} \exp\left(\frac{2t}{\sqrt{n}}X\right)\right)^n$  gives

$$\left(\mathbb{E} \exp\left[\frac{2t}{\sqrt{n}}X\right]\right)^n =$$

$$e^{-t^2/2} \left[ \left(1 - \frac{t^4/4}{2n}\right) + \left(1 - \frac{t^2/2(1-t^2/2)}{2n}\right) \left[ \frac{\delta}{b} \frac{(2t)^3}{n^{3/2}} + \frac{\tau}{24} \frac{(2t)^7}{n} \right] \right.$$

$$\left. + \left(1 - \frac{t^2/2(2-t^2/2)}{2n}\right) \frac{(n-1)}{n^2} \frac{\delta^2 (2t)^6}{72} \right] + o(n^{-1})$$

*throw out  $\frac{1}{n^2} \frac{\delta^3 (2t)^6}{42}$*

$$= e^{-t^2/2} \left[ 1 - \frac{t^4}{8n} + \frac{\delta}{b} \frac{(2t)^3}{n^{3/2}} + \frac{\tau}{24} \frac{(2t)^7}{n} + \frac{1}{n} \frac{\delta^2 (2t)^6}{72} \right] + o(n^{-1})$$

□

$$= e^{-t^2/2} \left[ 1 + \frac{\delta}{6} \frac{(zt)^3}{\sqrt{n}} + \frac{(\tau-3)}{24n} (zt)^4 + \frac{\delta^2 (zt)^6}{72n} \right] + o\left(\frac{1}{n}\right).$$

Remember, this is a Taylor expansion of  $\psi_{S_n}$ . Let

$$\tilde{\psi}_{S_n}(t) = e^{-t^2/2} \left[ 1 + \frac{\delta}{6} \frac{(zt)^3}{\sqrt{n}} + \frac{(\tau-3)}{24n} (zt)^4 + \frac{\delta^2 (zt)^6}{72n} \right]$$

Now, to get an approximation to the cdf of  $S_n$ , we

(i) invert  $\tilde{\psi}_{S_n}$  to get the corresponding pdf  $\tilde{f}_{S_n}$

(ii) take the antiderivative of  $\tilde{f}_{S_n}$ .

Inversion of  $\tilde{\psi}_{S_n}(t)$ :

$$\tilde{f}_{S_n}(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-itx} \tilde{\psi}_{S_n}(t) dt$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-itx} e^{-t^2/2} \left[ 1 + \frac{\delta}{6} \frac{(zt)^3}{\sqrt{n}} + \frac{(\tau-3)}{24n} (zt)^4 + \frac{\delta^2 (zt)^6}{72n} \right] dt$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-itx} e^{-t^2/2} dt + \frac{\delta}{6\sqrt{n}} \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-itx} e^{-t^2/2} (zt)^3 dt$$

$$+ \frac{(\tau-3)}{24n} \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-itx} e^{-t^2/2} (zt)^4 dt$$

$$+ \frac{\delta^2}{72n} \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-itx} e^{-t^2/2} (zt)^6 dt$$

$$= \phi(x) + \frac{\delta}{6\sqrt{n}} H_3(x) \phi(x) + \frac{\tau-3}{24n} H_4(x) \phi(x)$$

□

$$+ \frac{\gamma^2}{72n} H_6(x) \phi(x).$$

So we have

$$\tilde{f}_{S_n}(x) = \phi(x) \left[ 1 + \frac{\gamma}{6\sqrt{n}} H_3(x) + \frac{(\tau-3)}{24n} H_4(x) + \frac{\gamma^2}{72n} H_6(x) \right].$$

Now take the antiderivative of  $\tilde{f}_{S_n}(x)$ , making use of the fact

$$\frac{d}{dx} [H_k(x) \phi(x)] = -H_{k+1}(x) \phi(x).$$

$$\begin{aligned} \tilde{F}_{S_n}(x) &= \Phi(x) - \frac{\gamma}{6\sqrt{n}} H_2(x) \phi(x) - \frac{(\tau-3)}{24n} H_3(x) \phi(x) - \frac{\gamma^2}{72n} H_5(x) \phi(x) \\ &= \Phi(x) - \phi(x) \left[ \frac{\gamma}{6\sqrt{n}} H_2(x) + \frac{(\tau-3)}{24n} H_3(x) + \frac{\gamma^2}{72n} H_5(x) \right] \\ &= \Phi(x) - \phi(x) \left[ \frac{\gamma}{6\sqrt{n}} (x^2-1) + \frac{(\tau-3)}{24n} (x^3-3x) + \frac{\gamma^2}{72n} \overbrace{(x^5-10x^3+15x)}^{H_5(x)} \right]. \end{aligned}$$

This completes the derivation of the Edgeworth Expansions.

Note that

$$\gamma = \mathbb{E}X_1^3 = \mathbb{E}\left(\frac{X_1-0}{1}\right)^3 = \frac{\mu_3}{\sigma^3},$$

where  $\mu_3 = \mathbb{E}(X_1 - \mathbb{E}X_1)^3$  and  $\sigma^2 = \mathbb{E}(X_1 - \mathbb{E}X_1)^2$ , and

$$\tau-3 = \mathbb{E}\left(\frac{X_1-0}{1}\right)^4 = \frac{\mu_4}{\sigma^4} - 3,$$

where  $\mu_4 = \mathbb{E}(X_1 - \mathbb{E}X_1)^4$ .

□

We can also find Edgeworth expansions for the studentized pivot

$$\frac{\sqrt{n}(\bar{X}_n - \mu)}{\hat{\sigma}_n}$$

We state these results without derivation (see Hill (1992) for details).

Theorem (1<sup>st</sup> and 2<sup>nd</sup> order Edgeworth Expansions for studentized pivot):

Let  $Y_1, \dots, Y_n$  be iid with  $\mathbb{E} Y = \mu$ ,  $\text{Var} Y = \sigma^2 \in (0, \infty)$ ,  $\mathbb{E}|Y|^3 < \infty$ ,  $\mathbb{E}|Y|^4 < \infty$ ,  
and suppose

$$|\mathbb{E} \exp(itY)| < 2 \quad \text{for all } t \neq 0. \quad (\text{Non-lattice})$$

Then

$$\sup_{x \in \mathbb{R}} \left| \mathbb{P}\left(\frac{\sqrt{n}(\bar{Y}_n - \mu)}{\hat{\sigma}_n} \leq x\right) - \tilde{\Phi}_{n,3}(x) \right| = o(n^{-1/2})$$

as  $n \rightarrow \infty$  and

$$\sup_{x \in \mathbb{R}} \left| \mathbb{P}\left(\frac{\sqrt{n}(\bar{Y}_n - \mu)}{\hat{\sigma}_n} \leq x\right) - \tilde{\Phi}_{n,4}(x) \right| = o(n^{-1}),$$

where

$$\tilde{\Phi}_{n,3}(x) = \Phi(x) + \frac{1}{6\sqrt{n}} \frac{\mu_3}{\sigma^3} (2x^2 + 1) \phi(x)$$

$$\begin{aligned} \tilde{\Phi}_{n,4}(x) = \Phi(x) + \left\{ \frac{1}{6\sqrt{n}} \frac{\mu_3}{\sigma^3} (2x^2 + 1) + \frac{1}{12n} \cdot \left( \frac{\mu_4}{\sigma^4} - 3 \right) (x^3 - 3x) \right. \\ \left. - \frac{1}{18n} \cdot \frac{\mu_5}{\sigma^6} (x^5 + 2x^3 - 3x) - \frac{1}{4n} (x^3 + 3x) \right\} \phi(x). \end{aligned}$$

## EDGEWORTH EXPANSION FOR THE BOOTSTRAP

Consider again the three pivots from the previous lecture

$$\sqrt{n}(\bar{X}_n - \mu), \quad \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma}, \quad \text{and} \quad \frac{\sqrt{n}(\bar{X}_n - \mu)}{\hat{\sigma}_n},$$

where  $X_1, \dots, X_n$  are iid with  $\mathbb{E}X_i = \mu$ ,  $\text{Var} X_i = \sigma^2$ ,  $\bar{X}_n = n^{-1} \sum_{i=1}^n X_i$ ,  $\hat{\sigma}_n^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2$   
and the corresponding cdfs

$$G_{n,U}(x) = \mathbb{P}\left(\sqrt{n}(\bar{X}_n - \mu) \leq x\right)$$

$$G_n(x) = \mathbb{P}\left(\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \leq x\right)$$

$$G_{n,S}(x) = \mathbb{P}\left(\frac{\sqrt{n}(\bar{X}_n - \mu)}{\hat{\sigma}_n} \leq x\right) \quad \text{for all } x \in \mathbb{R}.$$

The Edgeworth expansion results, applied to these pivots, give

$$\sup_{x \in \mathbb{R}} \left| G_{n,U}(x) - \Phi\left(\frac{x}{\sigma}\right) \right| = O(n^{-1/2})$$

$$\sup_{x \in \mathbb{R}} \left| G_n(x) - \Phi(x) \right| = O(n^{-1/2})$$

$$\sup_{x \in \mathbb{R}} \left| G_{n,S}(x) - \Phi(x) \right| = O(n^{-1/2})$$

as  $n \rightarrow \infty$ .

We interpret these results by saying that the "accuracy" of the Normal approximation to  $G_{n,U}$ ,  $G_n$ , and  $G_{n,S}$  is of the order  $O(n^{-1/2})$ .

Now we present a result for the accuracy of the bootstrap estimators of the cdfs  $G_{n,U}$ ,  $G_n$ , and  $G_{n,S}$ .

Result: For  $X_1, \dots, X_n$  iid with mean  $\mu$ , variance  $\sigma^2 \in (0, \infty)$ ,  $\mathbb{E} X_i^3 = \mu_3$  and  $\mathbb{E} X_i^4 = \mu_4 < \infty$ , provided

$$|\mathbb{E} \exp(-tX_i)| < 1 \quad \text{for all } t \neq 0,$$

we have

$$(i) \quad \sup_{x \in \mathbb{R}} |\hat{G}_{n,U}(x) - G_{n,U}(x)| = O_p(n^{-1/2})$$

$$(ii) \quad \sup_{x \in \mathbb{R}} |\hat{G}_n(x) - G_n(x)| = o_p(n^{-1/2})$$

$$(iii) \quad \sup_{x \in \mathbb{R}} |\hat{G}_{n,S}(x) - G_{n,S}(x)| = o_p(n^{-1/2})$$

as  $n \rightarrow \infty$ .

Note, the accuracy  $o_p(n^{-1/2})$  is better than  $O_p(n^{-1/2})$  and  $O(n^{-1/2})$ .

We refer to the property of this improved accuracy as second-order correctness.

The studentized bootstrap will, by the above result, perform better than the unstandardized bootstrap.

Sketch of Proof for Result:

For (i), we have, by the Edgeworth expansion result

$$\sup_{x \in \mathbb{R}} \left| G_{n,U}(x) - \left\{ \Phi\left(\frac{x}{\sigma}\right) - \frac{1}{6\sqrt{n}} \frac{\mu_3}{\sigma^3} \left(\left(\frac{x}{\sigma}\right)^2 - 1\right) \phi\left(\frac{x}{\sigma}\right) \right\} \right| = O(n^{-1})$$

and it can be shown that

$$\sup_{x \in \mathbb{R}} \left| \hat{G}_{n,U}(x) - \left\{ \Phi\left(\frac{x}{\hat{\sigma}_n}\right) - \frac{1}{6\sqrt{n}} \frac{\hat{\mu}_{3,n}}{\hat{\sigma}_n^3} \left(\left(\frac{x}{\hat{\sigma}_n}\right)^2 - 1\right) \phi\left(\frac{x}{\hat{\sigma}_n}\right) \right\} \right| = O_p(n^{-1})$$

as  $n \rightarrow \infty$ , where  $\hat{\mu}_{3,n} = n^{-1} \sum_{i=1}^n (X_i - \bar{X}_n)^3$ .



From the above, we have

$$\begin{aligned} \sup_{x \in \mathbb{R}} \left| \hat{G}_{n,u}(x) - G_{n,u}(x) \right| &\leq O_p(n^{-1}) + \sup_{x \in \mathbb{R}} \left| \Phi\left(\frac{x}{\hat{\sigma}_n}\right) - \Phi\left(\frac{x}{\sigma}\right) \right| \\ &\quad + \frac{1}{6\sqrt{n}} \frac{\hat{\mu}_{3,n}}{\hat{\sigma}_n^3} \sup_{x \in \mathbb{R}} \left| \left[ \left(\frac{x}{\hat{\sigma}_n}\right)^2 - 1 \right] \phi\left(\frac{x}{\hat{\sigma}_n}\right) \right| \\ &\quad + \frac{1}{6\sqrt{n}} \frac{\mu_3}{\sigma} \sup_{x \in \mathbb{R}} \left| \left[ \left(\frac{x}{\sigma}\right)^2 - 1 \right] \phi\left(\frac{x}{\sigma}\right) \right| \\ &= O_p(n^{-1/2}), \end{aligned}$$

as  $n \rightarrow \infty$ , where the last equality requires a little more work.

For (ii), we have

$$\sup_{x \in \mathbb{R}} \left| G_n(x) - \left\{ \Phi(x) - \frac{1}{6\sqrt{n}} \frac{\mu_3}{\sigma^3} (x^2 - 1) \phi(x) \right\} \right| = O(n^{-1})$$

and it can be shown that

$$\sup_{x \in \mathbb{R}} \left| \hat{G}_n(x) - \left\{ \Phi(x) - \frac{1}{6\sqrt{n}} \frac{\hat{\mu}_{3,n}}{\hat{\sigma}_n^3} (x^2 - 1) \phi(x) \right\} \right| = O_p(n^{-1})$$

as  $n \rightarrow \infty$ .

Then we may write

$$\begin{aligned} \sup_{x \in \mathbb{R}} \left| \hat{G}_n(x) - G_n(x) \right| &\leq O_p(n^{-1}) + \frac{1}{6\sqrt{n}} \left| \frac{\hat{\mu}_{3,n}}{\hat{\sigma}_n^3} - \frac{\mu_3}{\sigma^3} \right| \sup_{x \in \mathbb{R}} \left| (x^2 - 1) \phi(x) \right| \\ &= o_p(n^{-1/2}), \end{aligned}$$

since  $\hat{\mu}_{3,n} \rightarrow \mu_3$  w.p. 1 and  $\hat{\sigma}_n^3 \rightarrow \sigma^3$  w.p. 1.

For (iii), we have

$$\sup_{x \in \mathbb{R}} \left| G_{n,S}(x) - \left\{ \Phi(x) + \frac{1}{6\sqrt{n}} \frac{\mu_3}{\sigma^3} (2x^2+1) \phi(x) \right\} \right| = O(n^{-1})$$

and it can be shown that

$$\sup_{x \in \mathbb{R}} \left| \hat{G}_{n,S}(x) - \left\{ \Phi(x) + \frac{1}{6\sqrt{n}} \frac{\hat{\mu}_{3,n}}{\hat{\sigma}_n^3} (2x^2+1) \phi(x) \right\} \right| = O_p(n^{-1})$$

as  $n \rightarrow \infty$ .

Then we have

$$\begin{aligned} \sup_{x \in \mathbb{R}} \left| \hat{G}_{n,S}(x) - G_{n,S}(x) \right| &\leq O_p(n^{-1}) + \frac{1}{6\sqrt{n}} \left| \frac{\hat{\mu}_{3,n}}{\hat{\sigma}_n^3} - \frac{\mu_3}{\sigma^3} \right| \sup_{x \in \mathbb{R}} \left| [2x^2+1] \phi(x) \right| \\ &= o_p(n^{-1/2}), \end{aligned}$$

using again  $\hat{\mu}_{3,n} \rightarrow \mu_3$  u.p.1 and  $\hat{\sigma}_n^3 \rightarrow \sigma^3$  u.p.1.